CSC 480/580 Principles of Machine Learning

05 Practical Considerations

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Outline

• The role of features in supervised learning

Classification metrics beyond error rate

Model Evaluation & Comparison

Debugging Learning Algorithms

• Bias / Variance Tradeoff

The role of features in supervised learning

The role of features in supervised learning

- Importance of feature representation
- Feature selection
- Feature transformation

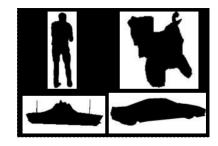
The importance of good feature representation

- Pixel representation:
 - represent an image as a w*h*3 dimensional vector
 - treat all coordinates in the same role
 - throw away all locality information in the image





- Shape representation:
 - represent a colored image with a w*h black-white image

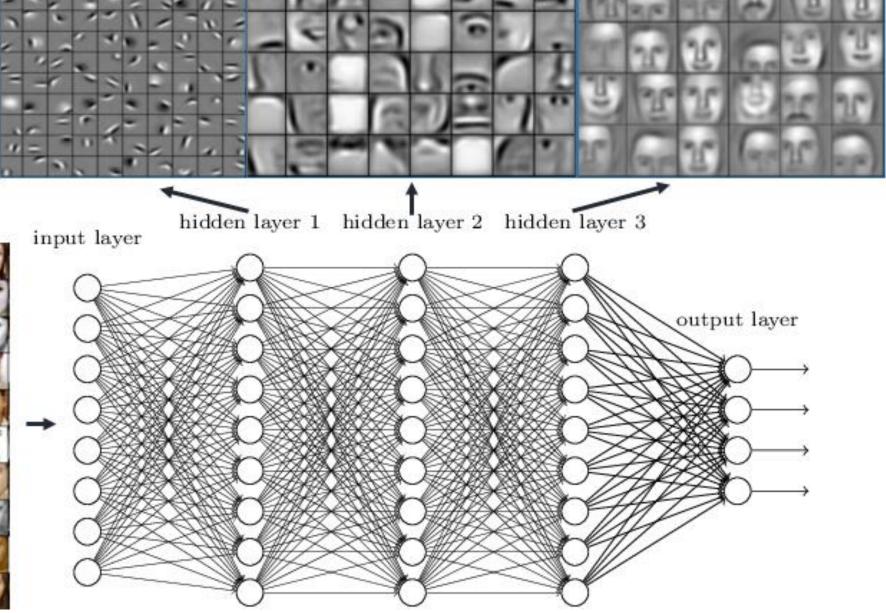


• Bag-of-words representation:

| | free | offer | lecture | cs | Spam? |
|---------|------|-------|---------|----|-------|
| Email 1 | 2 | 1 | 0 | 0 | +1 |
| Email 2 | 0 | 1 | 3 | 1 | -1 |

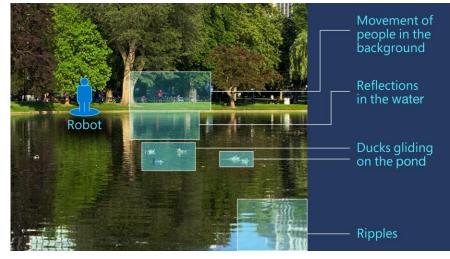
Deep neural networks learn hierarchical feature representations





Irrelevant and redundant features

- Irrelevant features
 - y is independent of f
 - y = Road walkability, f = duck activities in the pond



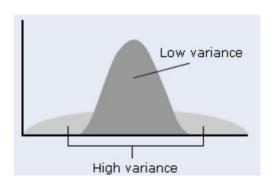
- If #features is large and #examples is small ⇒ spurious correlation between some feature & label
- Redundant features
 - Given f_1 , y is (nearly) independent of f_2
 - E.g. detecting fire hydrants
- Learning decision trees implicitly handles these two issues
- How about nearest neighbors / Perceptron?

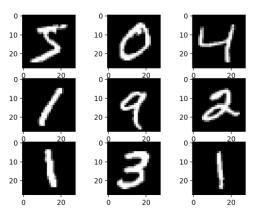


Feature pruning

- Removing features that are not very useful for prediction
 - E.g. text classification with bag-of-word representation, remove words that appear in <= K docs
 - E.g. digit classification, remove pixels with low variance

$$\mu_f = \frac{1}{N} \sum_{i=1}^{N} x_{i,f}$$
 $\sigma_f^2 = \frac{1}{N} \sum_{i=1}^{N} (x_{i,f} - \mu_f)^2$





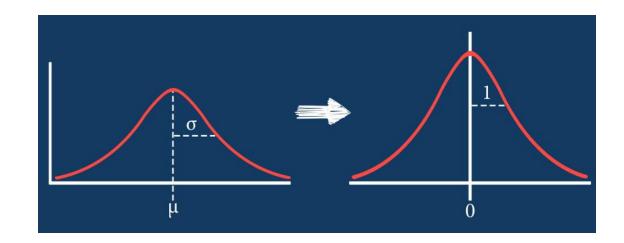
Feature normalization

• Centering:

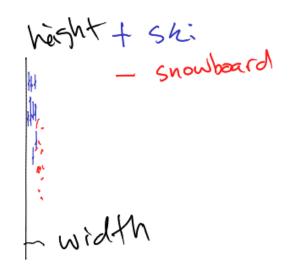
•
$$x'_{i,f} = x_{i,f} - \mu_f \Rightarrow \mu'_f = 0$$

Variance scaling:

•
$$x'_{i,f} = x_{i,f}/\sigma_f \Rightarrow (\sigma'_f)^2 = 1$$



- Absolute scaling
 - $x'_{i,f} = x_{i,f}/r_f$, where $r_f = \max_i |x_{i,f}| \Rightarrow$ range of $x'_{i,f}$ in [-1,+1]
- Same transformation applied to both training set and test data
- Aside: example normalization: $x_i' = \frac{x_i}{\|x_i\|}$ sometimes also can be applied



The role of features in supervised learning

- Importance of feature representation
- Feature selection
- Feature transformation

Example: Prostate Cancer Dataset

| | | _ | |
|-----------|--------|--------|-------|
| Term | LS | Ridge | Lasso |
| Intercept | 2.465 | 2.452 | 2.468 |
| lcavol | 0.680 | 0.420 | 0.533 |
| lweight | 0.263 | 0.238 | 0.169 |
| age | -0.141 | -0.046 | : |
| lbph | 0.210 | 0.162 | 0.002 |
| svi | 0.305 | 0.227 | 0.094 |
| lcp | -0.288 | 0.000 | |
| gleason | -0.021 | 0.040 | |
| pgg45 | 0.267 | 0.133 | |
| | | | |

Best LASSO model learns to ignore several features (age, lcp, gleason, pgg45).

Wait...Is **age** really not a significant predictor of prostate cancer? What's going on here?

Age is highly correlated with other factors and thus *not significant* in the presence of those factors

Best-Subset Feature Selection

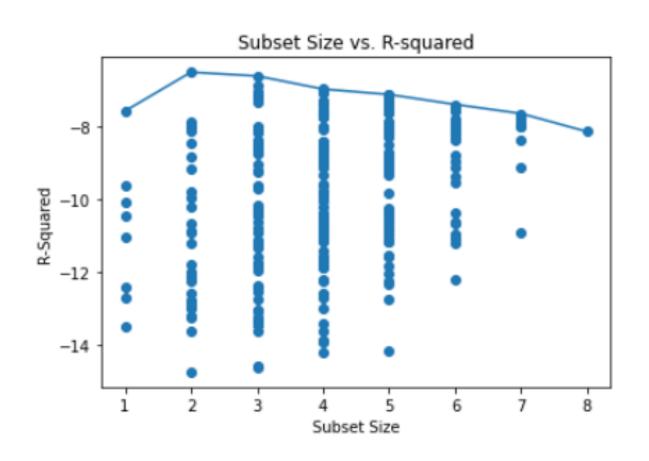
The optimal strategy for p features looks at models over *all possible combinations* of features,

```
For k in 1,...,p:
subset = Compute all subset of k-features (p-choose-k)

For kfeat in subset:
model = Train model on kfeat features
score = Evaluate model using cross-validation

Choose the model with best cross-validation score
```

Best-Subset Feature Selection: Prostate Cancer Dataset



Each marker is the cross-val R² score of a trained model for a subset of features

Data have 8 features, there are 8-choose-k subsets for each k=1,...,8 for a total of 255 models

Using 10-fold cross-val requires 10 x 255 = 2,550 training runs!

Feature Selection: Prostate Cancer Dataset

Best subset has lowest test error (lowest variance) with just 2 features

| LS | Best Subset | Ridge | Lasso |
|--------|----------------------------------------------------------------------------------------------------|------------------------------------------------------|------------------------------------------------------|
| 2.465 | 2.477 | 2.452 | 2.468 |
| 0.680 | 0.740 | 0.420 | 0.533 |
| 0.263 | 0.316 | 0.238 | 0.169 |
| -0.141 | | -0.046 | |
| 0.210 | | 0.162 | 0.002 |
| 0.305 | | 0.227 | 0.094 |
| -0.288 | | 0.000 | |
| -0.021 | | 0.040 | |
| 0.267 | | 0.133 | |
| 0.521 | 0.492 | 0.492 | 0.479 |
| 0.179 | 0.143 | 0.165 | 0.164 |
| | $2.465 \\ 0.680 \\ 0.263 \\ -0.141 \\ 0.210 \\ 0.305 \\ -0.288 \\ -0.021 \\ 0.267 \\ \hline 0.521$ | $\begin{array}{cccccccccccccccccccccccccccccccccccc$ | $\begin{array}{cccccccccccccccccccccccccccccccccccc$ |

[Source: Hastie et al. (2001)]

Forward Sequential Selection

An efficient method adds the most predictive feature one-by-one

```
featSel = empty
featUnsel = All features
For iter in 1,...,p:
 For kfeat in featUnsel:
   thisFeat = featSel + kfeat
   model = Train model on thisFeat features
   score = Evaluate model using cross-validation
 featSel = featSel + best scoring feature
 featUnsel = featUnsel - best scoring feature
Choose the model with best cross-validation score
```

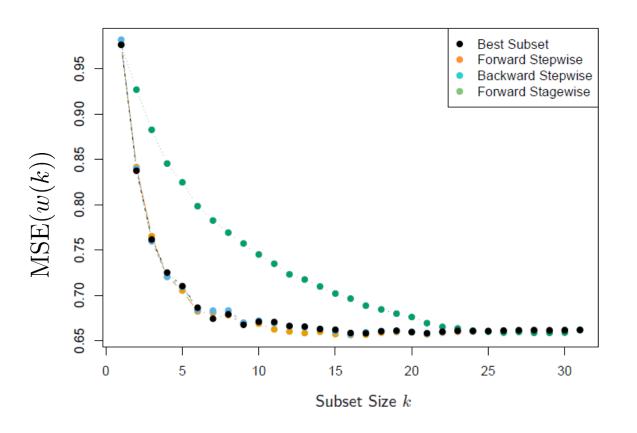
Backward Sequential Selection

Backwards approach starts with all features and removes one-by-one

```
featSel = All features
For iter in 1,...,p:
 For kfeat in featSel:
   thisFeat = featSel - kfeat
   model = Train model on thisFeat features
   score = Evaluate model using cross-validation
 featSel = featSel - worst scoring feature
Choose the model with best cross-validation score
```

Comparing Feature Selection Methods

Sequential selection is greedy, but often performs well...



Example Feature selection on synthetic model with p=30 features with pairwise correlations (0.85). True feature weights are all zero except for 10 features, with weights drawn from N(0,6.25).

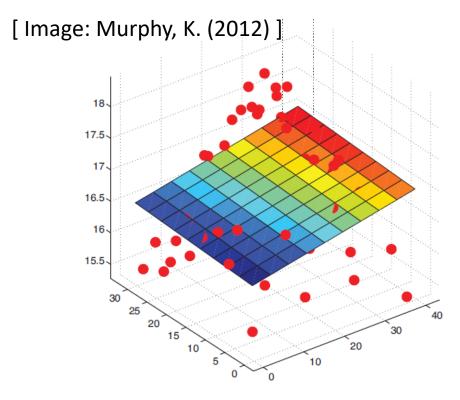
Sequential selection with p features takes $O(p^2)$ time, compared to exponential time for best subset

Sequential feature selection available in Scikit-Learn under: feature selection. Sequential Feature Selector

The role of features in supervised learning

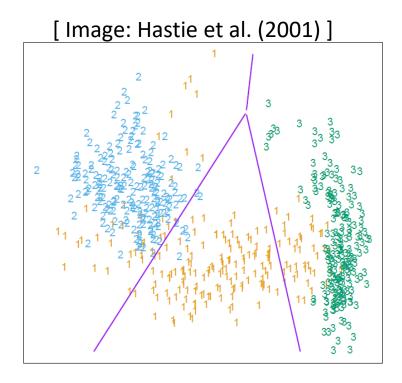
- Importance of feature representation
- Feature selection
- Feature transformation

Linear Models



Linear Regression Fit a *linear function* to the data,

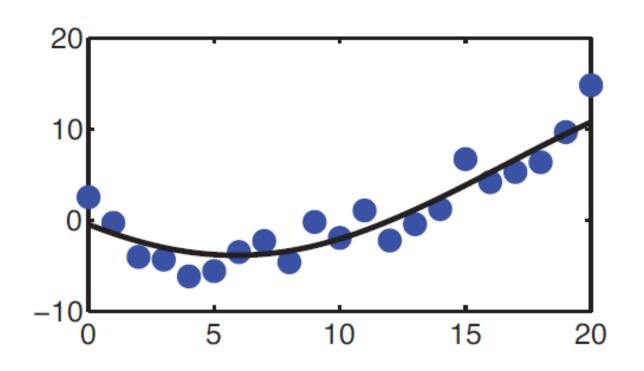
$$y = w^T x + b$$



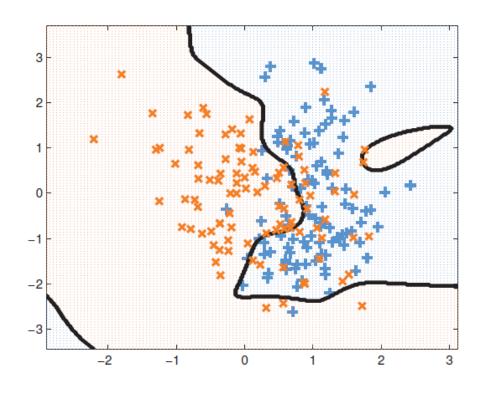
Linear Classification Learn a decision boundary that is *linear in the data*

$$y = sign(w^T x + b)$$

Nonlinear Data



What if our data are *not* well-described by a linear function?



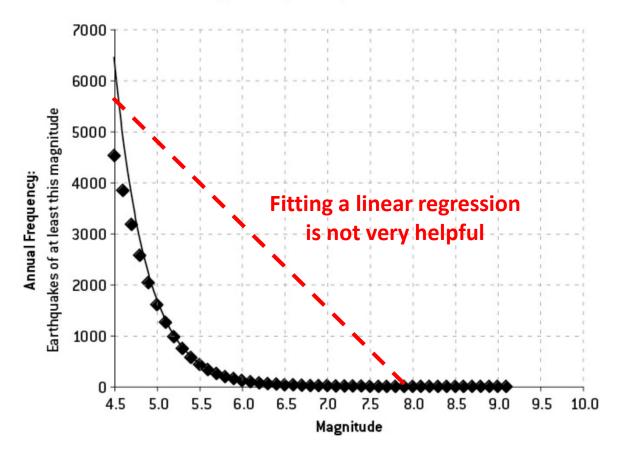
What if classes are not linearly-separable?

[Source: Murphy, K. (2012)]

Example: Earthquake Prediction

Suppose that we want to predict the number of earthquakes that occur of a certain magnitude. Our data are given by,



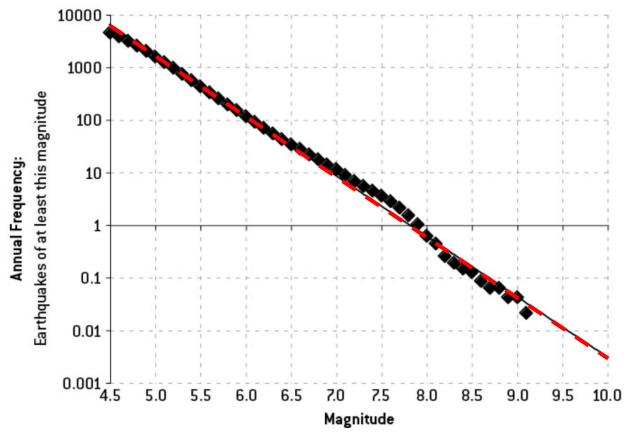


[Source: Silver, N. (2012)]

Example: Earthquake Prediction

Suppose that we want to predict the number of earthquakes that occur of a certain magnitude. Our data are given by,

FIGURE 5-3B: WORLDWIDE EARTHQUAKE FREQUENCIES, JANUARY 1964—MARCH 2012, LOGARITHMIC SCALE



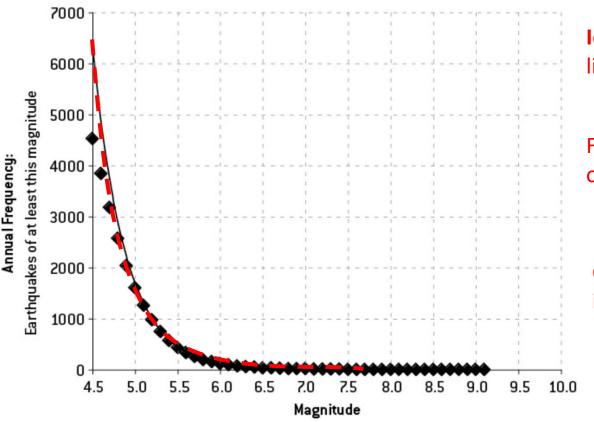
But plotting outputs on a logarithmic scale reveals a strong linear relationship...

[Source: Silver, N. (2012)]

Example: Earthquake Prediction

Suppose that we want to predict the number of earthquakes that occur of a certain magnitude. Our data are given by,





Idea Instead of fitting ordinary linear regression,

$$y = w^T x$$

First take the logarithm of output values y,

$$y = w^T \log(x)$$

Or take the exponential of input values x,

$$y = w^T \exp(x)$$

[Source: Silver, N. (2012)]

Basis Functions

- A basis function can be any function of the input features X
- Define a set of m basis functions $\phi_1(x), \ldots, \phi_m(x)$
- Fit a linear regression model in terms of basis functions,

$$y = \sum_{i=1}^{m} w_i \phi_i(x) = w^T \phi(x)$$

- Regression model is linear in the basis transformations
- Model is nonlinear in the data X

Common "All-Purpose" Basis Functions

Linear basis functions recover the original linear model,

$$\phi_m(x) = x_m$$

Returns mth dimension of X

- Quadratic. $\phi_m(x)=x_j^2$ or $\phi_m(x)=x_jx_k$ capture 2nd order interactions
- An order p polynomial $\phi \to x_d, x_d^2, \dots, x_d^p$ captures higher-order nonlinearities (but requires $O(d^p)$ parameters)
- Nonlinear transformation of single inputs,

$$\phi \to (\log(x_j), \sqrt{x_j}, \ldots)$$

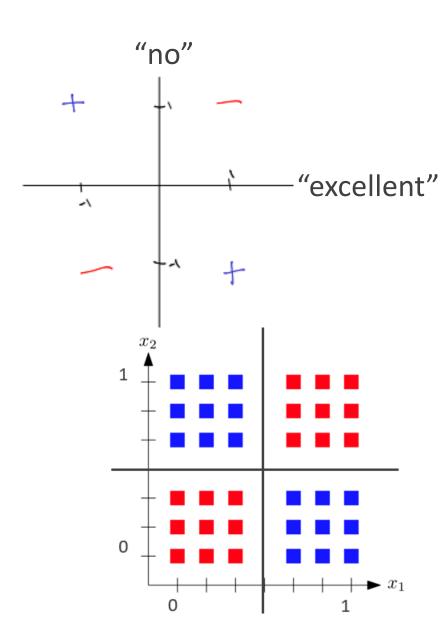
An indicator function specifies a region of the input,

$$\phi_m(x) = I(L_m \le x_k < U_m)$$

Feature transformations

- Combining features into a "meta-feature", e.g. $x_{no} \cdot x_{excellent}$
 - Useful for e.g. Perceptron learners
- In general, $\binom{d}{k}$ meta-features if allowed to combine k features
- Computationally cheaper alternative:
 - train a decision tree, use the meta-feature induced by leaves

- Logarithmic feature transformation
 - $x'_f \leftarrow \log_2(x_f)$ ("excellent" word count: 1->2 vs. 10->11)
 - $x_f' \leftarrow \log_2(x_f + 1)$



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• Classification metrics beyond error rate

Model Evaluation & Comparison

Debugging Learning Algorithms

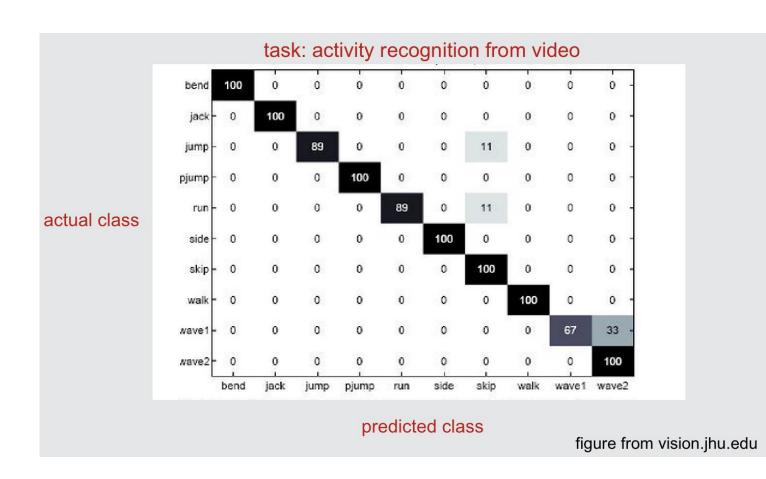
Bias / Variance Tradeoff

Classification metrics beyond error rate

Confusion matrix

• E.g. activity recognition

• $P(\hat{y} = \text{skip} \mid y = \text{jump}) = 11\%$



Class imbalance problem

- E.g., 5% pos, 95% negative.
- Implicit assumption: misclassifying positive example is more costly than misclassifying negative examples



- Tend to produce classifiers that always predict negative
- 2 alternatives:
 - Duplicate the minority class to make the positive and negative class balanced repeat every positive example w times, where w = P(y = -1)/P(y = +1)
 - Importance weighted classification: minimize $\sum_{i=1}^{n} w_i \ I(h(x_i) \neq y_i)$, where $w_i = 1$ when $y_i = -1$, $w_i = w$ when $y_i = +1$

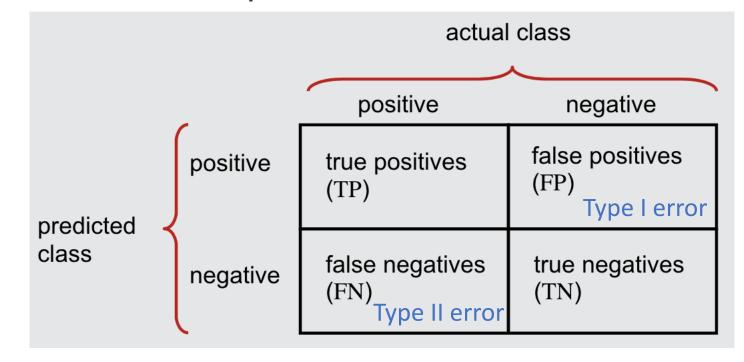
New measures of classification performance

True positive rate (TPR)

$$= \frac{\text{TP}}{\text{P}} = \frac{P(\hat{y}=+1,y=+1)}{P(y=+1)}$$

(aka recall, sensitivity)

- True negative rate (TNR) = $\frac{TN}{N}$ (specificity)
- False positive rate (FPR) = $\frac{FP}{N}$
- False negative rate (FNR) = $\frac{FN}{P}$



$$P = TP + FN$$

$$P = TP + FN$$
 $N = FP + TN$

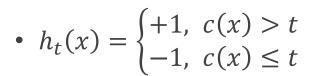
Applications:

- Search engine: precision & recall
- Cancer classification: FNR vs. FPR

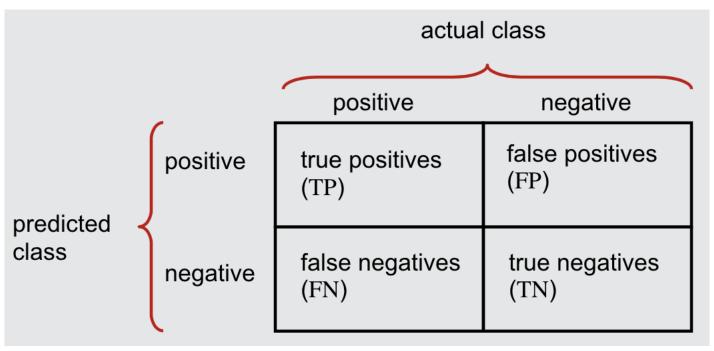
Adjusting TP, FP, TN, FN via thresholding

• Decision values (classification scores)

| $c(x_i)$ | y_i | |
|----------|-------|--|
| .99 | + | |
| .98 | + | |
| .72 | - | |
| .51 | - | |
| .24 | + | |



- Choice of threshold *t*:
 - $t = +\infty$: $h_t \equiv -1 \Rightarrow TPR = 0$, FPR = 0
 - t = 0: $h_t \equiv +1 \Rightarrow TPR = 1$, FPR = 1

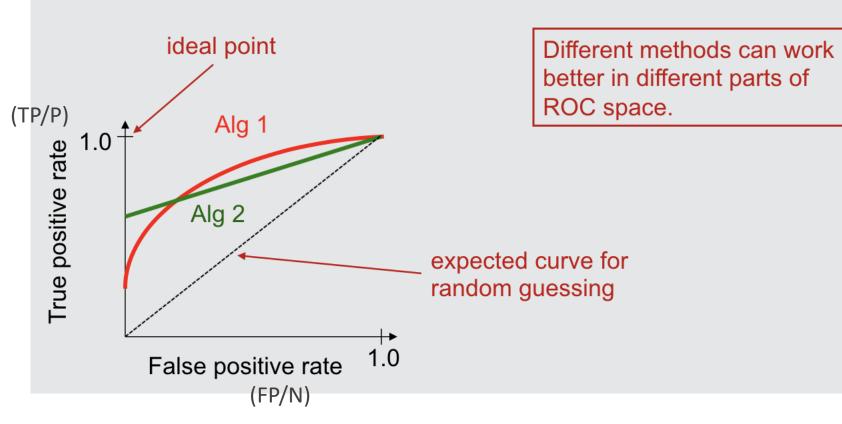


$$P = TP + FN$$
 $N = FP + FN$

$$TPR = \frac{TP}{P}, FPR = \frac{FP}{N}$$

ROC curve

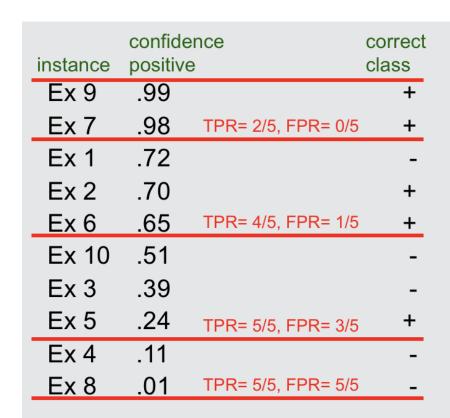
A Receiver Operating Characteristic (ROC) curve plots the TP-rate vs. the FP-rate as a threshold on the confidence of an instance being positive is varied

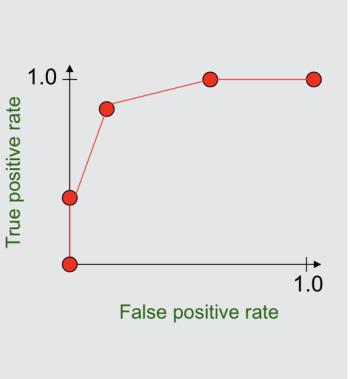


| $c(x_i)$ | y_i |
|----------|-------|
| .99 | + |
| .98 | + |
| .72 | + |
| .51 | - |
| .24 | - |

ROC curve

- Conceptually, consider every possible threshold, put a dot for each, and connect them.
- Actually, just need to care about when the 'correct class' changes
 - results in staircase shape, but diagonal line can still happen.
- A popular alternative: just plot when going from + to -. (what's shown here)





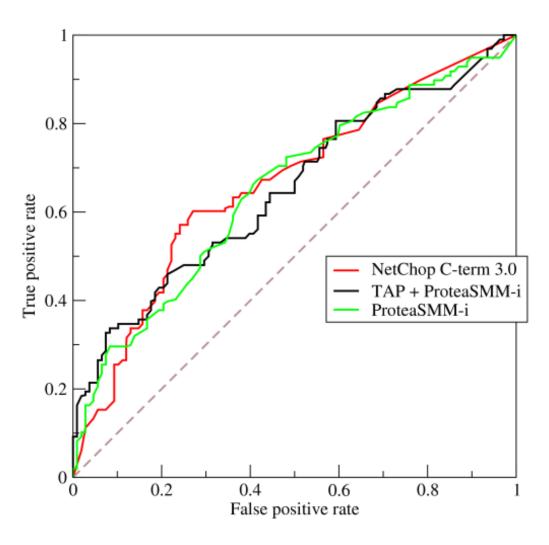


Calculating ROC curve

```
let (y^{(1)}, c^{(1)}) \dots (y^{(m)}, c^{(m)}) be the test-set instances sorted according to predicted confidence c^{(i)} that each instance is positive
let num_neg, num_pos be the number of negative/positive instances in the test set
TP = 0, FP = 0
last TP = 0
for i = 1 to m
    // find thresholds where there is a pos instance on high side, neg instance on low side
    if (i > 1) and (c^{(i)} \ne c^{(i-1)}) and (y^{(i)} = = \text{neg}) and (TP > last\_TP)
            FPR = FP / num\_neg, TPR = TP / num\_pos
           output (FPR, TPR) coordinate
           last TP = TP
    if y^{(i)} == pos
            ++TP
    else
           ++FP
FPR = FP / num\_neg, TPR = TP / num\_pos
output (FPR, TPR) coordinate
```

| | confider | nce | correct |
|----------|----------|--------------------|---------|
| instance | positive | | class |
| Ex 9 | .99 | | + |
| Ex 7 | .98 | TPR= 2/5, FPR= 0/5 | + |
| Ex 1 | .72 | | - |
| Ex 2 | .70 | | + |
| Ex 6 | .65 | TPR= 4/5, FPR= 1/5 | + |
| Ex 10 | .51 | | - |
| Ex 3 | .39 | | - |

ROC curve examples



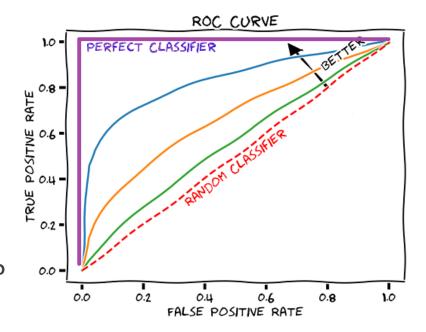
from Wikipedia

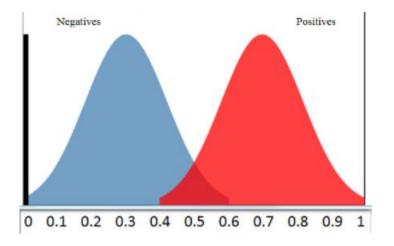
Area under ROC curve

- The boss says "could you just give me one number?"
- AUC: Area Under the ROC curve
 - Perfect scorer: AUC = 1
 - Random scorer: AUC = 0.5

- Quiz: Why does the perfect scorer c have a ROC that is ' Γ ' shaped ?
 - What would a ``worst'' ROC look like?

• Interpretation: "how well does c distinguish between + and -?"





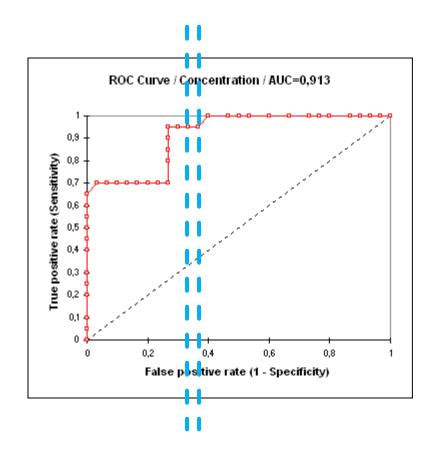
Area under ROC curve

• **Fact:** AUC of scorer *c* has the following formula:

$$AUC(c) = \frac{\sum_{(x_{-},-1)\in S_{-}} \sum_{(x_{+},+1)\in S_{+}} I(c(x_{+})>c(x_{-}))}{N_{-}\cdot N_{+}}$$

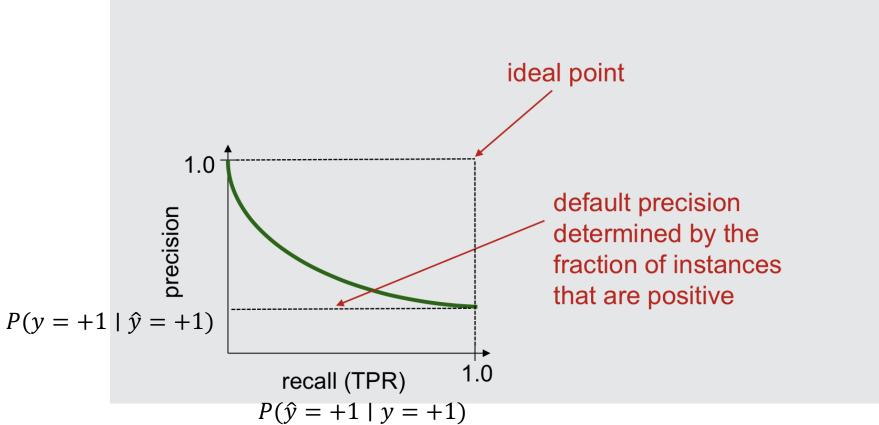
• Idea: the vertical slice corresponding to x_{-} has area

$$\frac{1}{N_{-}} \cdot \frac{\sum_{(x_{+},+1) \in S_{+}} I(c(x_{+}) > c(x_{-}))}{N_{+}}$$



Precision-Recall (PR) curve

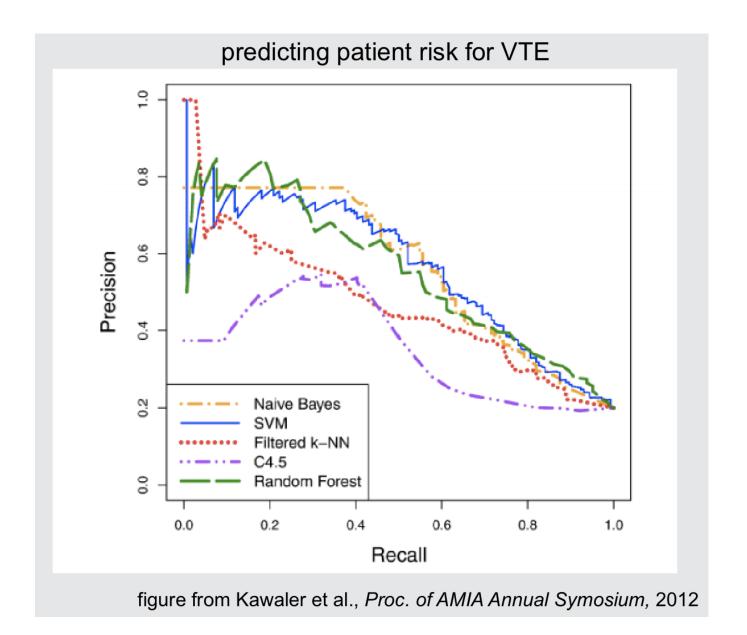
A precision/recall curve plots the precision vs. recall (TP-rate) as a threshold on the confidence of an instance being positive is varied



| $c(x_i)$ | y_i |
|----------|-------|
| .99 | + |
| .98 | + |
| .72 | + |
| .51 | - |
| .24 | - |

• This is *usually* a trade-off curve (not always): $t \downarrow \Rightarrow$ recall \(\frac{1}{2}\), precision usually \(\frac{1}{2}\)

PR-curve example



Summary of precision-recall

- Reporting one number
- Take the harmonic mean: **F1 score**

• Fact: minimum of two numbers <= harmonic mean <= geometric mean <= arithmetic mean

$$F_1 = rac{2}{ ext{recall}^{-1} + ext{precision}^{-1}}$$

- Emphasizes the smaller measure
 - E.g. recall = 0.1, precision = $0.9 \Rightarrow F_1 = 0.18$

Area under PR-curve is also a popular metric

| | 0.0 | 0.2 | 0.4 | 0.6 | 0.8 | 1.0 |
|-----|------|------|------|------|------|------|
| 0.0 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |
| 0.2 | 0.00 | 0.20 | 0.26 | 0.30 | 0.32 | 0.33 |
| 0.4 | 0.00 | 0.26 | 0.40 | 0.48 | 0.53 | 0.57 |
| 0.6 | 0.00 | 0.30 | 0.48 | 0.60 | 0.68 | 0.74 |
| 0.8 | 0.00 | 0.32 | 0.53 | 0.68 | 0.80 | 0.88 |
| 1.0 | 0.00 | 0.33 | 0.57 | 0.74 | 0.88 | 1.00 |

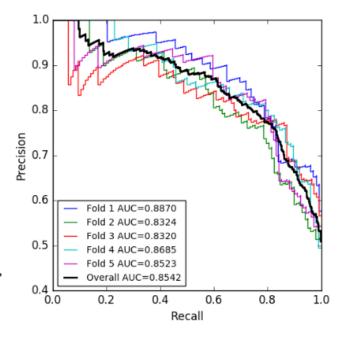
Table 5.2: Table of f-measures when varying precision and recall values.

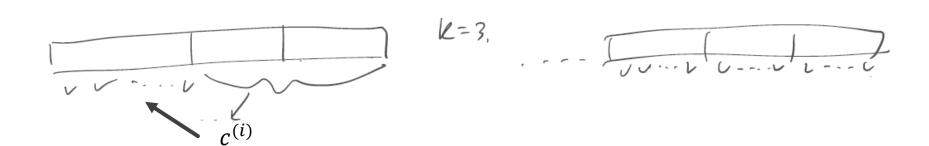
How to plot ROC/PR curve when training set is small?

- k-fold cross-validation (CV):
 - Obtain k curves and plot them all



• **Pooled prediction:** compute c(x) for all training examples using k-fold CV.





Model Evaluation & Comparison

Motivation: evaluating & comparing ML models

Example

- Your ML model *f* has test set error = 6.9%
- Your nemesis, Gabe's, ML model g has test set error = 6.8%
- How confident are we to conclude that g has smaller population error than that of f?
- Intuition: We should be more (less) confident, if the test set is larger (smaller)
- Our uncertainty can be quantified with a confidence interval
- Determining the best model can be done rigorously with hypothesis testing

Disclaimer: we only focus on the key ideas (standard stats courses spend >= 5 lectures on this)

Confidence Intervals (CIs)

Intuition Find an interval such that we are pretty sure it encompasses the

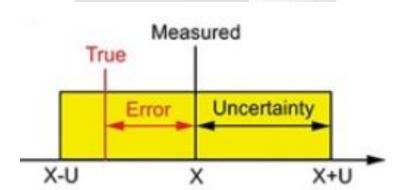
true parameter value (e.g. model accuracy).

Also commonly used when making measurements

Given data X_1, \ldots, X_n and confidence $\alpha \in (0,1)$ find interval (a,b) such that,

$$P(\theta \in (a,b)) \ge 1 - \alpha$$

In English the interval (a,b) contains the true parameter value θ with probability at least $1-\alpha$



- Allows us to quantify our *uncertainty* about θ (wider CI => higher uncertainty)
- Intervals must be computed from data $a(X_1,\ldots,X_n)$ and $b(X_1,\ldots,X_n)$
- Interval (a,b) is random, parameter θ is not random (it is fixed)
- Intuitively: α smaller => wider CI; n larger => narrower CI;

https://www.nelpretech.com/resources/measurement-uncertainty

Confidence interval: caveat in interpretation

- Suppose $\theta = \text{true width of the window (in cm)}$
- By making 100 independent measurements, we have computed a 95% confidence interval: [247.5, 252.5]
- Does that mean:

with probability 0.95,
$$\theta \in [247.5, 252.5]$$
 (*)

?

- <u>Hint</u>: What is the randomness in the above statement?
- **Answer:** The above interpretation is **incorrect** " $\theta \in [247.5, 252.5]$ " is either true or false there is no randomness in this statement
- Then: what does $P(\theta \in (a(X), b(X))) \ge 0.95$ truly mean?

Interpretation

On day 1, you collect data and construct a 95 percent confidence interval for a parameter θ_1 . On day 2, you collect new data and construct a 95 percent confidence interval for an unrelated parameter θ_2 . On day 3, you collect new data and construct a 95 percent confidence interval for an unrelated parameter θ_3 . You continue this way constructing confidence intervals for a sequence of unrelated parameters $\theta_1, \theta_2, \ldots$ Then 95 percent of your intervals will trap the true parameter value. There is no need to introduce the idea of repeating the same experiment over and over.

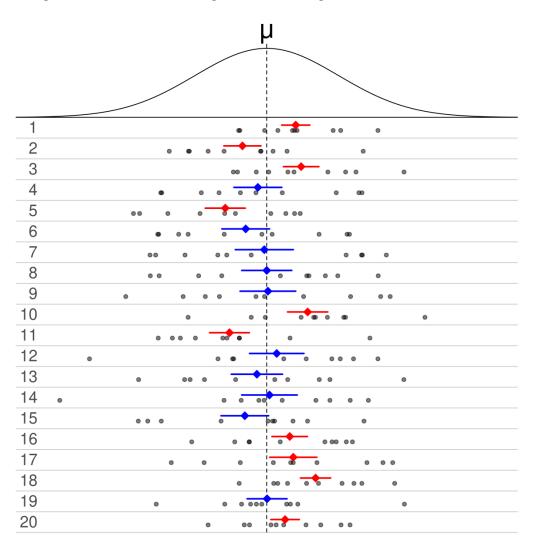
Another interpretation: as long as we collect a *representative* sample (which happens 95% of the time), its induced confidence interval contains θ

As long as we are not extremely unlucky, our induced confidence interval contains heta

[Source: Wasserman, L. 2004]

Knowledge Check

What is the confidence level of this confidence interval construction?



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Bias / Variance Tradeoff

Confidence Intervals (CIs)

Intuition Find an interval such that we are pretty sure it encompasses the

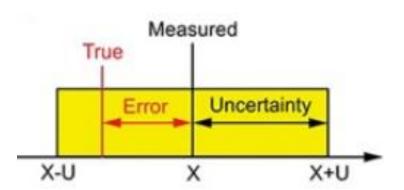
true parameter value (e.g. model accuracy).

Also commonly used when making measurements

Given data X_1, \ldots, X_n and confidence $\alpha \in (0,1)$ find interval (a,b) such that,

$$P(\theta \in (a,b)) \ge 1 - \alpha$$

In English the interval (a,b) contains the true parameter value θ with probability at least $1-\alpha$



- Allows us to quantify our *uncertainty* about θ (wider CI => higher uncertainty)
- Intervals must be computed from data $a(X_1,\ldots,X_n)$ and $b(X_1,\ldots,X_n)$
- Interval (a,b) is random, parameter θ is not random (it is fixed)
- Intuitively: α smaller => wider CI; n larger => narrower CI;

https://www.nelpretech.com/resourc es/measurement-uncertainty

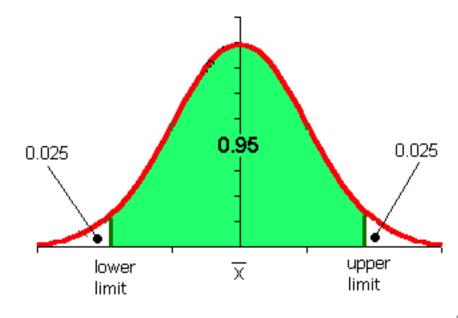
CI construction: a standard recipe

- Step 1: Construct an estimator for θ based on S -- call it $\widehat{\theta}_S$ (center of CI)
- Step 2: Let CI $I(S) \coloneqq [\hat{\theta}_S w, \hat{\theta}_S + w]$, where w is chosen such that for all θ , $P(\theta \in [\hat{\theta}_S w, \hat{\theta}_S + w]) \ge 1 \alpha$
 - Equivalently,

$$P(\theta - \hat{\theta}_S) \le w \ge 1 - \alpha$$

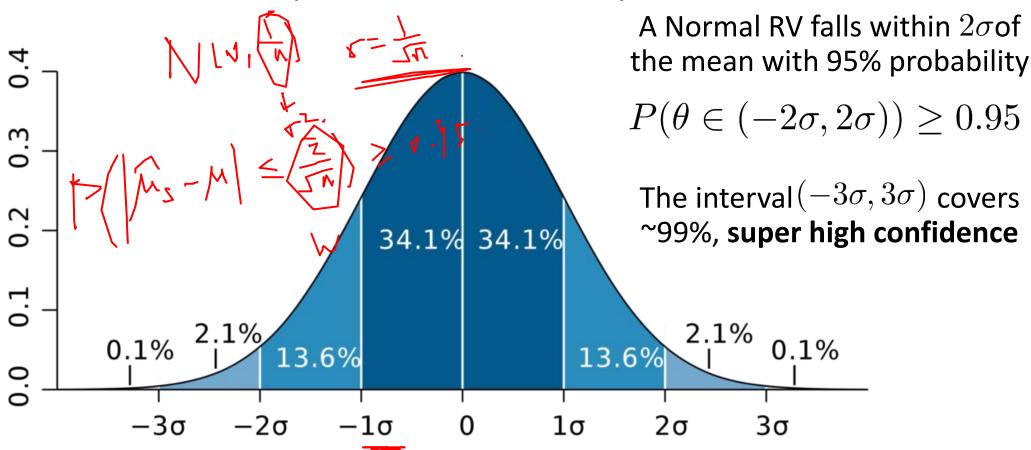
Important example: confidence interval for normal mean

- $D_{\mu} = N(\mu, 1), S = (X_1, ..., X_n) \sim D_{\mu}^n$
- Define $\hat{\mu}_S = \frac{1}{n} \sum_{i=1}^{n} X_i$ Known variance
- $\hat{\mu}_{S} \mu \neq N\left(0, \frac{1}{n}\right)$
- How to choose w such that $P(|\hat{\mu}_S \mu| \le w) \ge 1 \alpha$?



Confidence Intervals of the Normal Distribution

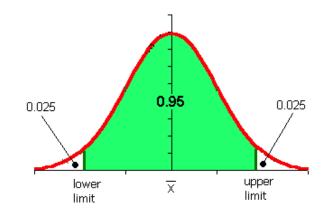
Many estimators follow a normal distribution with enough data (central limit theorem)



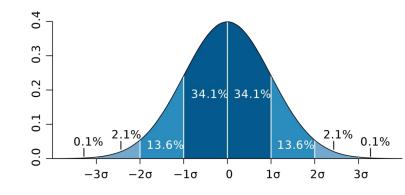
For various reasons, 95% has become standard confidence level

CI for normal mean (cont'd)

- $\hat{\mu}_S \mu \sim N\left(0, \frac{1}{n}\right)$
- How to choose w such that $P(|\hat{\mu}_S \mu| \le w) \ge 1 \alpha$?



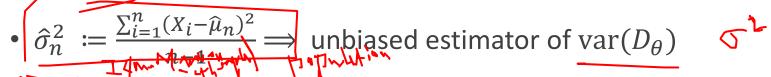
- Note: $Z = \sqrt{n} (\hat{\mu}_S \mu) \stackrel{\sim}{\sim} N(0,1)$ Central limit theorem
- Suffices to find z_{α} such that $P(|Z| \le z_{\alpha}) \ge 1 \alpha$, and let $w = \frac{z_{\alpha}}{\sqrt{n}}$



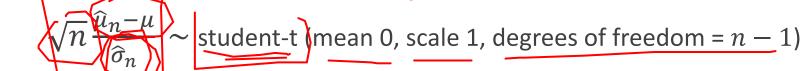
- Final (1α) -confidence interval construction for μ : $I(S) = \left[\hat{\mu}_S \frac{z_\alpha}{\sqrt{n}}, \hat{\mu}_S + \frac{z_\alpha}{\sqrt{n}}\right]$
- E.g. 95%-confidence interval for μ : $I(S) = \left[\hat{\mu}_S \frac{1.96}{\sqrt{n}}, \hat{\mu}_S + \frac{1.96}{\sqrt{n}}\right]$

CI for means of general distributions, unknown variance

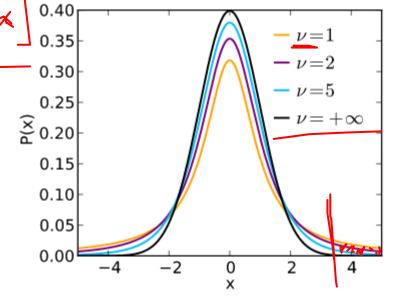
• Given D_{θ} with mean parameter θ with *unknown* variance



• Fact: Let $X_1, \dots, X_n \sim N(\mu, \sigma^2)$, and $\hat{\mu}_n \coloneqq \frac{1}{n} \sum_{i=1}^n X_i$, then

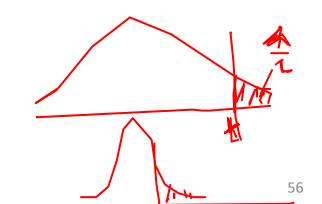


How do we estimate variance of algorithm performance?



import scipy.stats as st alpha = 0.05 st.t.ppf(1-alpha/2,df=2) => 4.302652729911275 st.t.ppf(1-alpha/2,df=5) => 2.5705818366147395 st.t.ppf(1-alpha/2,df=10) => 2.2281388519649385 st.t.ppf(1-alpha/2,df=30) => 2.0422724563012373 st.t.ppf(1-alpha/2,df=100)

=> 1.9839715184496334



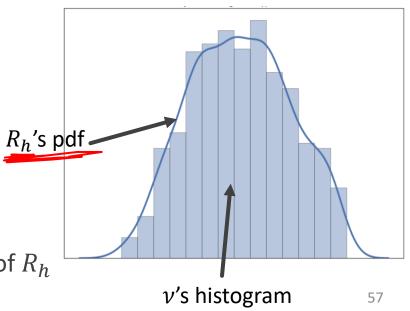
Bootstrapping Cl

- General approach for constructing CI for any statistic (e.g. F1 score, AUROC, FPR, ...)
- Goal: using sample S, estimate property h of D (:=h(D)) using confidence intervals
- **Step 1:** choose center of CI = h(S)
- Step 2: choose w_L , w_R such that $P(h(D) \in [h(S) w_L, h(S) w_R]) \ge 1 \alpha$

Equivalently:
$$P(h(S) - h(D) \in [w_R, w_L]) \ge 1 - \alpha$$

How?

- <u>Idea:</u> estimate the distribution of h(S) h(D), denoted by R_h by *bootstrapping* (resampling)
 - perform n times of "sampling with replacement" from S
 - repeat B times (e.g., B = 100) to obtain $S_1, ..., S_B$
 - take $\underline{v} := \underline{\text{empirical distribution}}$ of $\{h(S_b) h(S)\}_{b=1}^B$, as the 'shape' of R_h



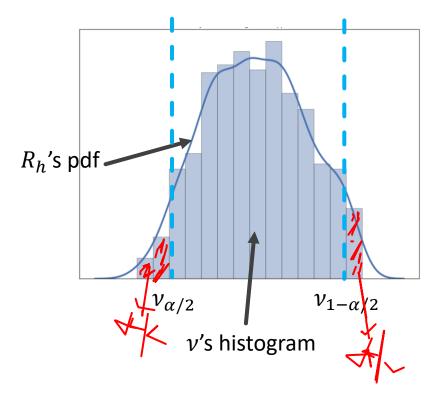
Bootstrapping CI: Implementation

- Goal: choose w_R, w_L , s.t. $P(h(S) h(D) \in [w_R, w_L]) \ge 1 \alpha$
- ν (histogram of $\{h(S_b) h(S)\}_{b=1}^B$) approximates R_h , the probability distribution of h(S) h(D)
- Choose:

$$w_R = \underline{v_{\alpha/2}}, w_L = v_{1-\alpha/2}$$
 -- quantiles of ν

Intuition: this well-approximates quantiles of R_h

- How to calculate ν 's quantiles?
 - Sort $\{h(S_b) h(S)\}_{b=1}^B$ in increasing order; say v[0..(B-1)]
 - $v_{1-\alpha/2} := \text{the top } 0.025 \text{ (i.e., v[int(0.975*B)])}$
 - $\nu_{\alpha/2} \coloneqq$ the bottom 0.025 (i.e., v[int(0.025*B)])
- Final bootstrap CI: $[h(S) w_L, h(S) w_R] = [h(S) v_{1-\alpha/2}, h(S) v_{\alpha/2}]$



Bootstrap Example

Example Suppose we have LSAT scores and GPA for 15 law students and wish to

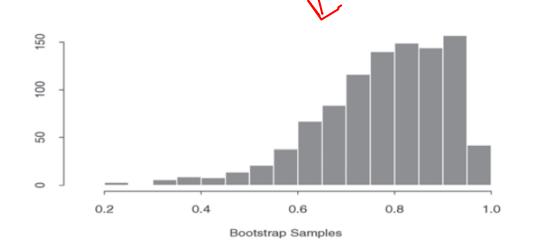
estimate the correlation between LSAT and GPA:

| LSAT | 576 651 | | | 580 572 | 661 |
|------|--------------|--|--|--------------|------|
| GPA | 3.39 3.36 | | | 3.07 2.88 | 3.43 |

PAS 3.5 3.6 Seo 580 600 620 640 660 LSAT

95% Bootstrap confidence interval from B=1000 estimates of the **correlation**,

$$.78 \pm .274 \Rightarrow (.51, 1.00)$$

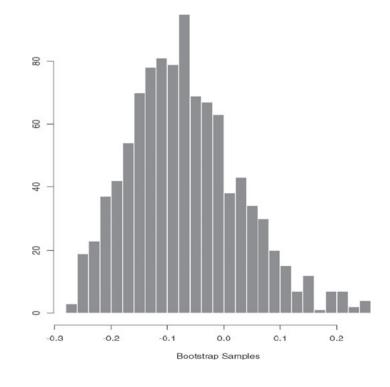


[Source: Wasserman, L. 2004]

Bootstrap Example

Eight subjects who used medical patches to infuse a hormone into the blood using three treatments: placebo, old-patch, new-patch

| subject | placebo | old | new | old - placebo | new - old |
|---------|---------|-------|-------|---------------|-----------|
| 1 | 9243 | 17649 | 16449 | 8406 | -1200 |
| 2 | 9671 | 12013 | 14614 | 2342 | 2601 |
| 3 | 11792 | 19979 | 17274 | 8187 | -2705 |
| 4 | 13357 | 21816 | 23798 | 8459 | 1982 |
| 5 | 9055 | 13850 | 12560 | 4795 | -1290 |
| 6 | 6290 | 9806 | 10157 | 3516 | 351 |
| 7 | 12412 | 17208 | 16570 | 4796 | -638 |
| 8 | 18806 | 29044 | 26325 | 10238 | -2719 |



Estimate whether relative efficacy is the same under new drug,

$$\theta = \frac{\mathbf{E}[\text{new} - \text{old}]}{\mathbf{E}[\text{old} - \text{placebo}]}$$

Bootstrap B=1,000 samples yields 95% confidence interval,

$$\theta \in (-0.24, 0.15)$$

Hypothesis testing: motivation

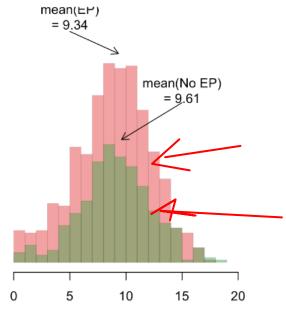
- How to prove that your new system A is better than existing one B
- **Ex (document retrieval):** A and B generate 10 lists of documents given 10 different keywords => the evaluator provides rating (1-5) for both lists

| Evaluator | 1 | 2 | 3 | 4 | 5 | 6 | ••• |
|-----------|-----|---|---|---|---|---|-----|
| А | 5 4 | 2 | 2 | 5 | 4 | 2 | ••• |
| В | 4 | 1 | 1 | 4 | 3 | 1 | |

• Ex (classification): 100 test data points => Record A and B's classification correctness in each

Two-sample hypothesis testing: definition

- Distribution family $\{D_{\theta}\}$ with parameter θ
- Samples $S_X = (X_1, ..., X_n)$ and $S_Y = (Y_1, ..., Y_n)$ drawn iid from distribution $D_{\underline{\theta_X}}$ and $D_{\underline{\theta_Y}}$, respectively
- Equality test version:
 - Null hypothesis H_0 : $\theta_X = \theta_Y$
 - Alternative hypothesis $H_1:\theta_X \neq \theta_Y$
- E.g. $D_{\mu} = N(\mu, 1), H_0: \mu_X = \mu_Y$
- Design hypothesis tester *T* such that:
 - T maps $S = (S_X, S_Y)$ to hypothesis index 0 or 1
 - Type-I error rate $P_{H_0}(\underline{T(S)} = 1) \le \alpha (\alpha : \text{significance level}; \text{say, 0.05})$
 - Intuition: do not overturn the null hypothesis unless we have strong evidence conservatism



Paired t-test

• $\underline{S_X} = (X_1, ..., X_n)$ and $\underline{S_Y} = (Y_1, ..., Y_n)$ drawn iid from distribution $D_{\theta_X} = N(\mu_X, \sigma_X^2)$ and $D_{\theta_Y} = N(\mu_Y, \sigma_Y^2)$, respectively

| • | H_0 : | μ_X | = | μ_Y |
|---|---------|---------|---|---------|
|---|---------|---------|---|---------|

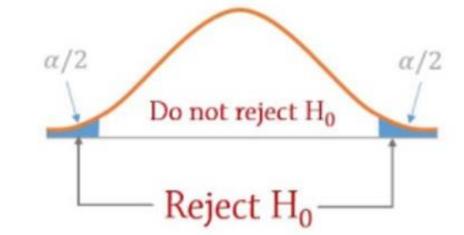
•
$$H_1$$
: $\mu_X \neq \mu_Y$

• Here, μ_X , μ_Y , σ_X^2 , σ_Y^2 are all unknown

| Evaluator | 1 | 2 | •• |
|----------------|---|---|----|
| <u>—</u> А | 5 | 2 | |
| — _В | 4 | 1 | |

An intuitive approach:

- Let $\delta_i := X_i Y_i$, for all i = 1, ..., n
- Let $\bar{\delta_n} \coloneqq \frac{1}{n} \sum_{i=1}^n \delta_i$

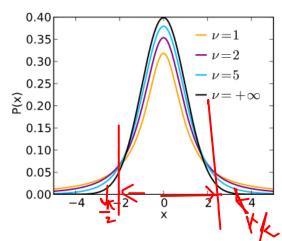


- Intuition: reasonable to accept H_0 if $|\bar{\delta}_n| \le c$, for some c > 0
- Design hypothesis test T so that $P_{H_0}(T(S) = 1) \le \alpha \Rightarrow c$ needs to be set large enough
- Challenge: $\bar{\delta}_n$'s distribution not known, since σ_X^2 , σ_Y^2 are unknown

Paired t-test

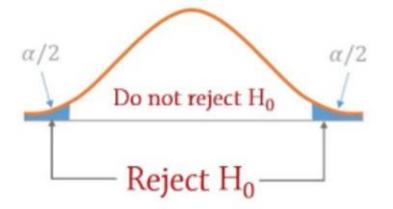
- Observe: Under H_0 , $\delta_i \sim N(0, \sigma^2)$, i = 1, ..., n, where $\sigma^2 = \sigma_X^2 + \sigma_Y^2$
- Recall <u>Fact</u>: Let $\delta_1, \dots, \delta_n \sim N(0, \underline{\sigma^2})$, and $\overline{\delta_n} \coloneqq \frac{1}{n} \sum_{i=1}^n \delta_i$, $\widehat{\sigma_n}^2 \coloneqq \frac{\sum_{i=1}^n (\delta_i \overline{\delta_n})^2}{n-1}$

$$Z = \sqrt{n} \frac{\overline{\delta}_n}{\widehat{\sigma}_n}$$
 ~ student-t (mean 0, scale 1, degrees of freedom = $n-1$)



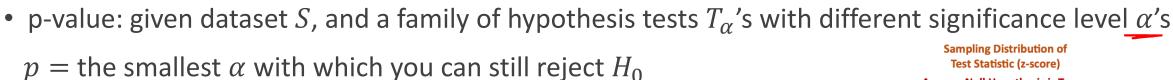
- Let's ask "under H_0 , what is a plausible range of values of Z with failure rate $\alpha=0.05$?"
 - Find the 0.025, 0.975-quantiles of $Z = \langle t_{0.025}, t_{0.975} \rangle$
 - Hypothesis tester

$$T(S) = I(Z \notin [t_{0.025}, t_{0.975}]) = I\left(\sqrt{n} \frac{\bar{\delta}_n}{\hat{\sigma}_n} \notin [t_{0.025}, t_{0.975}]\right)$$



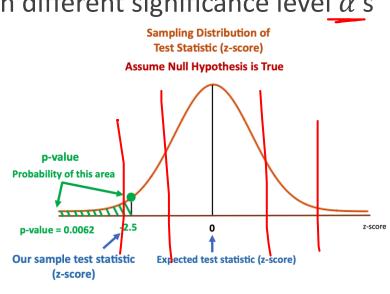
Hypothesis testing: additional remarks

- Confidence intervals can be used for hypothesis testing
 - $S = (X_1, ..., X_n)$ drawn iid from distribution D_{μ}
 - H_0 : $\mu = 0$
 - $H_1: \mu \neq 0$
 - I is a $(1-\alpha)$ -CI construction for μ => hypothesis test $T(S)=I(0\not\in I(S))$ has significance α



Smaller p-value => result more significant

- Hypotheses should be decided before observing data
 - Otherwise results may be misleading (p-hacking)



0.95

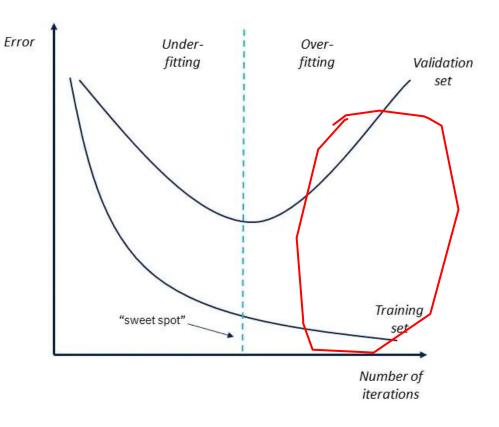
0.025

0.025

upper limit

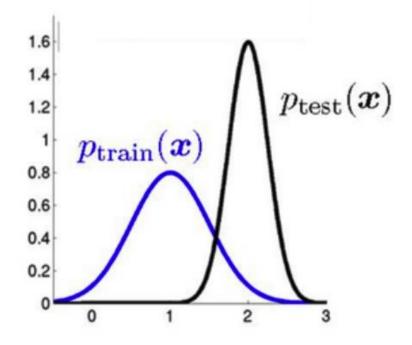
Is the problem with generalization to test data?

- Is it doing well on training?
 - Unrealistic to do better on test than on training
 - If not, problem may be *representation*: need better features or better data
- If it does well on training then problem is *generalization*
 - Model may be too complicated (overfitting)
 - Too many features, not enough training data



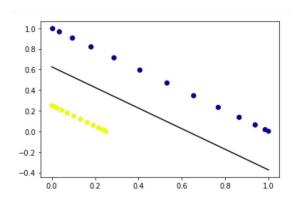
Is there a mismatch between training and test?

- Training data may be inadequate
- If enough training & test data:
 - Do results change with different train / test split?
 - If so then test distribution is probably strange



Is the learning algorithm implemented correctly?

- Is it optimizing the loss function that you intended?
- Try measuring / visualizing your loss function during training-is it going down?
- Do the data meet your algorithmic assumptions?
- Hand-craft datasets where you know the desired behavior
 - KNN on XOR function
 - Perceptron on data that is trivially linearly-separable
 - Decision tree on axis-aligned data
 - Generally, create dataset that meets assumptions of your algorithm



Do you have adequate representation?

- Your feature set could be inadequate
- For binary classification try this...
 - Add a feature (maybe call it CheatingIsFun)
 - Set value to +1 for positive instances and -1 for negative instances
 - This feature is a *perfect indicator*-problem is now trivially solvable
 - Does your algorithm solve it?
- If your algorithm doesn't get near 0% error then you may have a bug! (or more data / less features)
- If it does then you need better features or a different model (e.g. decision tree vs. linear model)

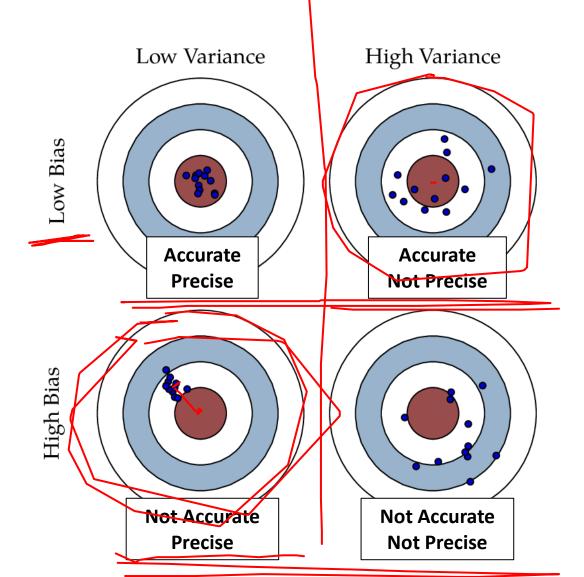
Do you have enough data?

- Conventional wisdom: Always have at least as many training data as you have learnable model parameters
 - Caveat: not required in modern overparameterized setting
- Try training on 80% of your training data
 - Does performance suffer?
 - How much? A lot?
 - If so then getting more data is likely helpful
 - If not then you may be data saturated-look elsewhere
- More training data should never lead to worse performance (just slower training)

Bias / Variance Tradeoff

Bias-Variance Tradeoff

Suppose an archer takes multiple shots at a target...



Bias-Variance Tradeoff

Is an unbiased estimator "better" than a biased one? It depends...

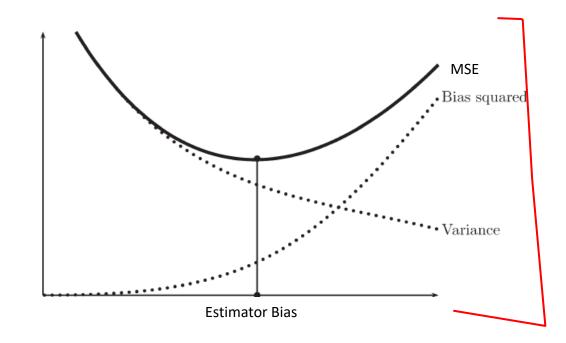
Evaluate the quality of estimate $\hat{\theta}$ using **mean squared error**,

$$MSE(\hat{\theta}) = \mathbf{E}\left[(\hat{\theta} - \theta)^2\right] = bias^2(\hat{\theta}) + \mathbf{Var}(\hat{\theta}) , bias(\hat{\theta}) = E[\hat{\theta}] - \theta$$

MSE for unbiased estimators is just,

$$MSE(\hat{\theta}) = Var(\hat{\theta})$$

- Bias-variance is fundamental tradeoff in statistical estimation
- MSE increases as square of bias
- Estimators with small bias (but low variance)
 can have lower MSE than unbiased estimators



Bias-Variance Decomposition

$$MSE(\hat{\theta}) = \mathbf{E} \left[(\hat{\theta}(X) - \theta)^{2} \right]$$

$$= \mathbf{E} \left[(\hat{\theta} - \mathbf{E}[\hat{\theta}] + \mathbf{E}[\hat{\theta}] - \theta)^{2} \right]$$

$$= \mathbf{E} [(\hat{\theta} - \mathbf{E}[\hat{\theta}])^{2}] + 2(\mathbf{E}[\hat{\theta}] - \theta)\mathbf{E}[\hat{\theta} - \mathbf{E}[\hat{\theta}]] + \mathbf{E}[(\mathbf{E}[\hat{\theta}] - \theta)^{2}]$$

$$= \mathbf{E} [(\hat{\theta} + \mathbf{E}[\hat{\theta}])^{2}] + (\mathbf{E}[\hat{\theta}] - \theta)^{2}$$

$$= Var(\hat{\theta}) + bias^{2}(\hat{\theta})$$

Other materials

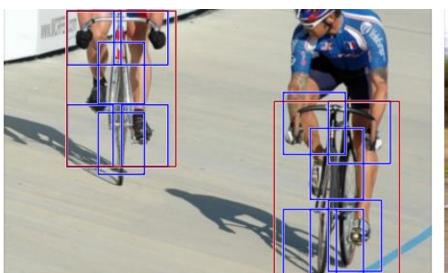
- Bootstrap confidence intervals: https://math.mit.edu/~dav/05.dir/class24-prep-a.pdf
- Permutation test: https://www.jwilber.me/permutationtest/
- STAT 566 lecture slides (at UA): https://www.math.arizona.edu/~jwatkins/stat566s20s.html

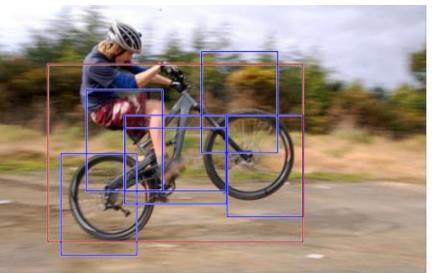
Next lecture (2/20)

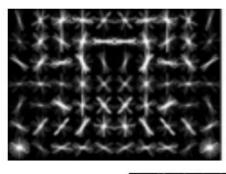
• Linear models revisited: classification, regression, loss minimization formulations

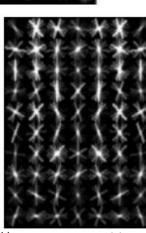
Assigned reading: CIML Chapter 7

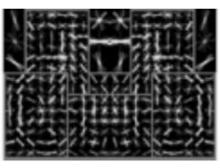
Note: We are skipping Chapter 6 for now!

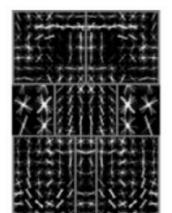


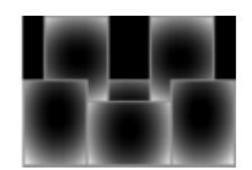


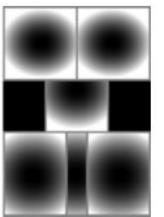






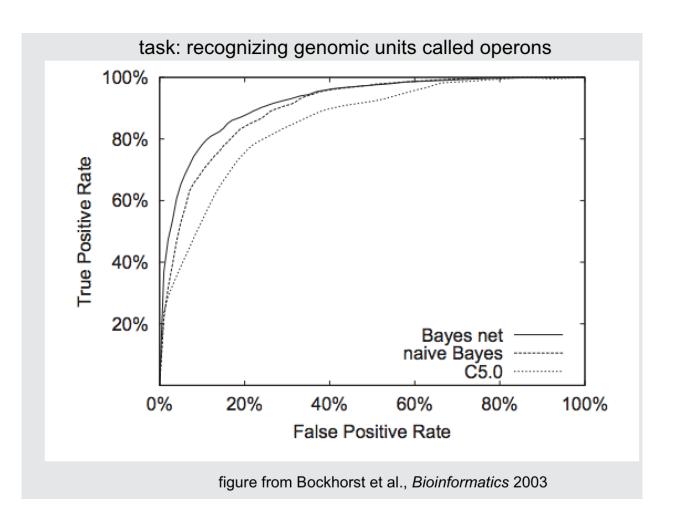


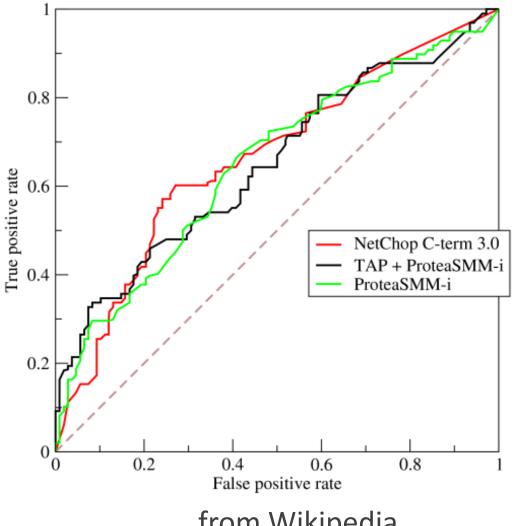




https://computervisionblog.wordpress.com/2016/01/23/distributed-code-or-grandmother-cells-insights-from-convolutional-neural-networks/

ROC curve examples





from Wikipedia

Warning

Question How should we interpret a confidence interval (e.g. 95%)?

$$P(\theta \in (a(X), b(X))) \ge 0.95$$

Hint Think about what is random and what is not...

Warning

Question How should we interpret a confidence interval (e.g. 95%)?

$$P(\theta \in (a(X), b(X))) \ge 0.95$$

Hint Think about what is random and what is not...

This is NOT a probability statement about θ .

Warning

Question How should we interpret a confidence interval (e.g. 95%)?

$$P(\theta \in (a(X), b(X))) \ge 0.95$$

Hint Think about what is random and what is not...

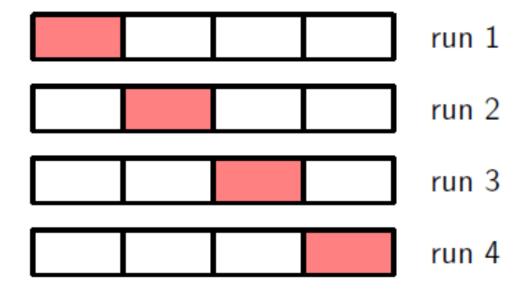
Confidence interval: caveat in interpretation

- Suppose θ = average height of people in Tucson (in cm)
- Using a sample of size 100, we have computed a 95% confidence interval: [168, 171]
- Does that mean:

with probability 0.95,
$$\theta \in [168, 171]$$
 (*

- **Hint**: What is the randomness in the above statement?
- **Answer:** The above interpretation is incorrect " $\theta \in [168, 171]$ " is either true or false there is no randomness in this statement
- Then: what does $P(\theta \in (a(X), b(X))) \ge 0.95$ truly mean?

Cross-Validation



K-fold Cross Validation Partition training data into K "chunks" and for each run select one chunk to be validation data

For each run, fit to training data (K-1 chunks) and measure accuracy on validation set. Average model error across all runs. Estimate variance.

Algorithm 8 CrossValidate(LearningAlgorithm, Data, K)

```
// store lowest error encountered so far
1: \hat{\epsilon} \leftarrow \infty
\hat{\alpha} \leftarrow \text{unknown}
                                           // store the hyperparameter setting that yielded it
_{3:} for all hyperparameter settings \alpha do
      err \leftarrow []
                                                 // keep track of the K-many error estimates
      for k = 1 to K do
         train \leftarrow \{(x_n, y_n) \in Data : n \mod K \neq k-1\}
         test \leftarrow \{(x_n, y_n) \in Data : n \mod K = k-1\} // test every Kth example
         model ← Run LearningAlgorithm on train
         err \leftarrow err \oplus error of model on test // add current error to list of errors
      end for
      avgErr \leftarrow \text{mean of set } err
                                                             Can also estimate variance here
      if avgErr < \hat{\epsilon} then
       \hat{\epsilon} \leftarrow avgErr
                                                                     // remember these settings
13:
         \hat{\alpha} \leftarrow \alpha
                                                              // because they're the best so far
      end if
16: end for
```

Drawback Need to perform training K times for each model.

Bootstrapping Cl

empirical distribution of $\{X_1, ..., X_n\}$: $\frac{1}{n} \sum_{i=1}^{n} \delta_{X_i} \text{ where } \delta_X \text{ is a dirac delta function}$

- Goal: estimate property h of D (:=h(D)) using confidence intervals, using sample S (e.g. h=F1 of model f)
- Idea: estimate the distribution of h(S)-h(D), denoted by R_h by bootstrapping (resampling)
 - perform n times of "sampling with replacement" from S
 - repeat B times (e.g., B = 100) to obtain $S_1, ..., S_B$
 - take $\nu \coloneqq \underline{\text{empirical distribution}}$ of $\{h(S_b) h(S)\}_{b=1}^B$, as the 'shape' of R_h



Quantile interval: sort values and take top/bottom-quantiles (see next slide)

